

Lecture 17: Intro to Graphical Models, VI

Def: (Undirected) graphical model w/ pairwise interactions:

Let $\{\psi_{ij}\}_{(ij) \in F}$ be compatibility functions $\{\pm 1\}^2 \rightarrow \mathbb{R}_{\geq 0}$ that dictate interactions b/t particles

The Gibbs measure is dist. over $\{\pm 1\}^n$ given by

$$M(x) \triangleq \frac{1}{Z} \prod_{(ij) \in F} \psi_{ij}(x_i, x_j), \quad Z \triangleq \sum_{x \in \{\pm 1\}^n} \prod_{(ij) \in F} \psi_{ij}(x_i, x_j)$$

"spins"

where Z is the partition function, i.e. normalizing constant.

We will use the shorthand

$$M \propto \prod_{(ij)} \psi_{ij}(x_i, x_j)$$

Example ("Ising model"):

$$\psi_{ij}(x_i, x_j) = \exp(-\beta A_{ij} x_i x_j), \text{ so}$$

$$M(x) \propto \exp\left(-\frac{\beta}{2} \underbrace{x^T A x}_{\text{"energy"}}\right)$$

for $A \in \mathbb{R}^{n \times n}$ a symmetric matrix with zero diagonal

β : "inverse temperature"

A : "Hamiltonian" / "interaction matrix"

$$E(x) \triangleq \frac{\beta}{2} x^T A x = -\lg \left(\prod_{(ij) \in F} \psi_{ij}(x_i, x_j) \right): \text{"energy"}$$

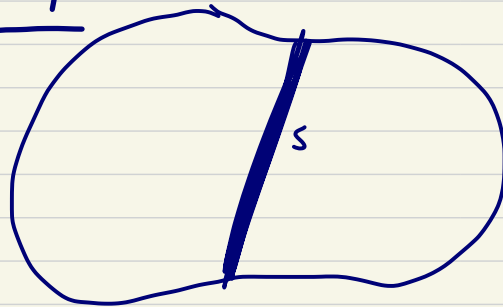
As $\beta \rightarrow 0$, $\mu \rightarrow \text{Unif}(\{\pm 1\}^n)$
 $\beta \rightarrow \infty$, $\mu \rightarrow \text{Unif}(\{\text{energy minimizers}\})$

Can think of A as adjacency matrix of weighted graph. Denote this by G .

$$\partial_i \triangleq \{j \text{ s.t. } A_{ij} \neq 0\} = \{j \text{ s.t. } (i,j) \in E\},$$

i.e. the neighbors of node i in G

Markov property:



If $[n] \setminus S$ decomposes into disjoint pieces, then marginal distributions on the pieces are independent, conditioned on any assignment to the spins on S

e.g. if we condition on x_{∂_i} , then conditional dist. on x_i is independent of conditional dist. on rest of the spins. For Ising model:

$$\Pr[x_i = \sigma \mid x_{\partial_i} = s] \propto \exp\left(\beta \sum_{j \in \partial_i} A_{ij} s_j \sigma\right)$$

* Negative sign is b/c of inconvenient culture clash b/t physics and CS: physicists want to minimize energy, in CS we want to maximize $x^T B x$, e.g. in MAXCUT

2 fundamental algorithmic tasks in inference:

① computing the partition function Z

② Sampling from Gibbs measure μ

Note alg. for ① \Rightarrow alg. for ② and vice versa
("equivalence of counting + sampling")

Challenge: Z is sum of exponentially many terms, so
in many cases we expect it is computationally hard to
compute...

e.g. if $\Psi_{ij}(x_i, x_j) = \mathbb{1}[x_i \neq x_j]$ for all $(i, j) \in E(G)$,

$Z = \#$ independent sets of G ("#P-complete",
i.e. very hard)

So our goal will be to approximate Z / approximately
sample from Gibbs measure μ

This
with
↓
☆

Some approaches:

- Markov chain Monte Carlo (MCMC)

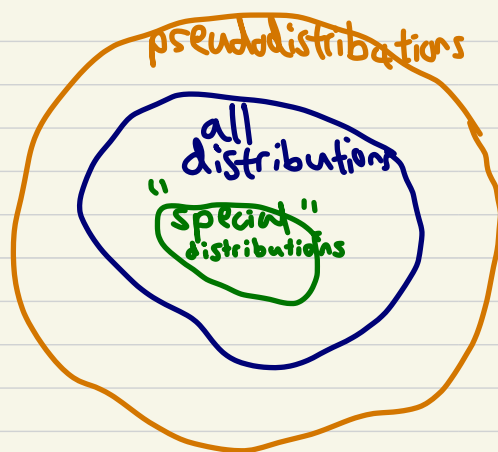
- Variational inference (VI)

- Diffusion models (very recent, more on this at the
end of the course)

VI: approximate μ by dist. from family ρ of "simpler" distributions that are easy to sample from (e.g. product distributions, aka naive mean-field):

$$\min_{\nu \in \rho} KL(\nu \parallel \mu) \quad (*)$$

Note: - if ρ is all distributions, minimizer is $\nu^* = \mu$
(Gibbs' inequality)
- "opposite" of SoS relaxation



An issue: can't even evaluate the objective function in (*), let alone optimize!

Fortunately, this particular issue is not really an issue:

$$KL(\nu \parallel \mu) = \sum_{\nu} \ln \frac{\nu}{\mu}$$

$$G[\nu] \triangleq -H(\nu) + \sum_{\nu} [\mathcal{E}]$$

"Gibbs free energy functional" /

-1 x "evidence lower bound" (ELBO)

(Note: Convex in ν)

$$= \sum_{\nu} \ln \frac{\nu}{e^{-\mathcal{E}}/Z}$$

free energy of the Gibbs measure

$$= \sum_{\nu} \ln \nu + \sum_{\nu} \mathcal{E} - \ln(\nu Z)$$

$\underbrace{\sum_{\nu} \ln \nu}_{\text{negative entropy } -H(\nu)}$
 $\underbrace{\sum_{\nu} \mathcal{E}}_{\text{average energy}}$
 $\underbrace{-\ln(\nu Z)}_{\text{independent of } \nu!}$

if ν simple, easy to evaluate

easy to approximate

so to minimize $KL(\nu \parallel \mu)$, suffices to minimize $G[\nu]$ which is easy to evaluate

Interpretation of G as "regularized energy":
for Ising model, recall $\mathcal{E}(x) = \frac{\beta}{2} x^T A x$, so

$$G[\nu] = \frac{\beta}{2} x^T A x - \underbrace{H(\nu)}_{\text{"entropy regularization"}}$$

"hot"
When β small, minimizer prioritizes maximizing entropy
 β big, minimizer prioritizes minimizing avg. energy
"cold"

$G[V]$ efficiently computable, but computationally intractable to optimize a priori...

Next lecture: powerful heuristic, belief propagation (BP), for solving $\min_V G[V]$.

2 interpretations of the heuristic:

① dynamic programming

② finding stationary points of a relaxation of the Gibbs free energy (Bethe free energy)
(see supplemental notes)

BP as dynamic programming (see slides for lecture 18)

Let's first shift focus to easier task than full-blown VI:
marginal estimation dist. μ_i over each node is a Bernoulli random variable, goal is to estimate it

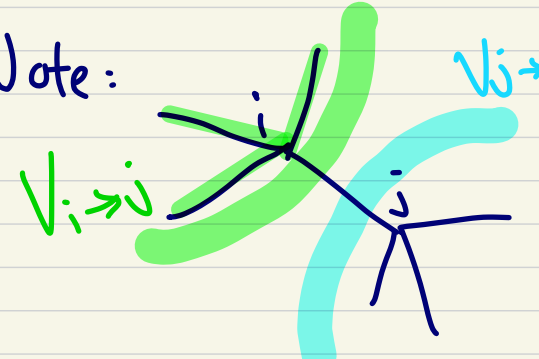
Physics motivation:


physicists care about limiting objects, and one important that they consider as $n \rightarrow \infty$ is the empirical dist. over marginals, i.e.


$$q_n(z) \stackrel{\delta}{=} \frac{1}{n} \sum_{i=1}^n \mathbb{1}[z = \mu_i]$$

and, given a sequence of Gibbs measures $(\mu^{(n)})$, want to understand $\lim_{n \rightarrow \infty} q_n$

To motivate the algorithm, assume G is a tree

Note:  $V_{j \rightarrow i}$ removing (i,j) from tree splits G into two subtrees

 $\bar{V}_{i \rightarrow j}$ ($V_{j \rightarrow i}$ + edge (i,j))

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To sample from μ_i ,

1). Sample spins on subtrees $V_{j \rightarrow i}$ for $j \in \partial_i$, yields assignment $s \in \{\pm 1\}^{\partial_i}$ to ∂_i

2). Sample from conditional dist. on x_i , i.e.

$$\mathbb{P}_i[x_i = \sigma \mid x_{\partial_i} = s] \propto \prod_{j \in \partial_i} \Psi_{ij}(\sigma, s_j)$$

By law of total probability,

$$\begin{aligned} \Pr[x_i = \sigma] &\propto \sum_{S \in \{\pm 1\}^{\mathcal{D}_i}} \prod_{j \in \mathcal{D}_i} \Pr[x_j = s_j] \psi_{ij}(\sigma, s_j) \\ &\stackrel{\text{b/c marginal dists on spins in } \mathcal{V}_{j \rightarrow i} \text{ are independent across } j\text{'s}}{=} \prod_{j \in \mathcal{D}_i} \underbrace{\sum_{s_j \in \{\pm 1\}} \Pr[x_j = s_j]}_{M_{\mathcal{V}_{j \rightarrow i}}} \psi_{ij}(\sigma, s_j) \end{aligned}$$

(~~✗✗~~)

(~~✗✗✗~~)

proportional to $\Pr[x_i = \sigma]$
 $M_{\mathcal{V}_{j \rightarrow i}}$

(i.e. can express marginals of $M_{\mathcal{V}_{j \rightarrow i}}$ in terms of marginals of $M_{\mathcal{V}_{j \rightarrow i}}$)

Unsatisfying b/c we've gone from $\Pr_{\mu}^{\mu}[x_i = \sigma]$ to $\Pr_{M_{\mathcal{V}_{j \rightarrow i}}}^{\mu}[x_i = \sigma]$, but we're very close.

Define messages :

$$m_{\sigma}^{j \rightarrow i} \stackrel{\Delta}{=} \Pr(x_j = \sigma)$$

$\mu_{V_{j \rightarrow i}}$

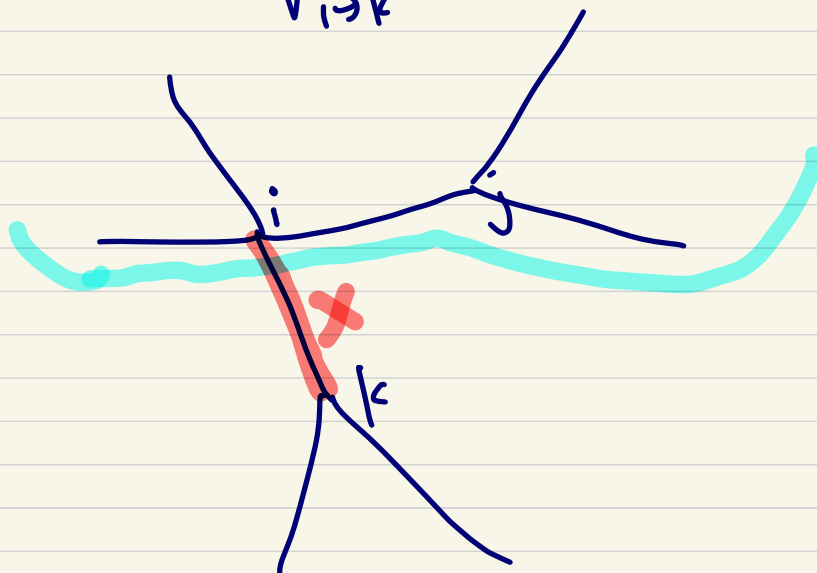
$$\overline{m}_{\sigma}^{j \rightarrow i} \stackrel{\Delta}{=} \Pr(x_i = \sigma)$$

$\mu_{\overline{V}_{j \rightarrow i}}$

Then ~~(*)~~ can be written as

$$\textcircled{I} \quad \overline{m}_{\sigma}^{j \rightarrow i} \propto \sum_{s \in \{\pm 1\}} m_s^{j \rightarrow i} \cdot \psi_{ij}(\sigma, s)$$

Also note that ~~(*)~~ can be modified to apply to $\mu_{V_{i \rightarrow k}}$ instead of μ , i.e.



previously, (604) gave

$$\Pr_{\mu} [x_i = \sigma] \propto \prod_{j \in \partial i} \bar{m}_{\sigma}^{j \rightarrow i}$$

after removing edge (i, k) , we get

$$\textcircled{\text{II}} \quad m_{\sigma}^{i \rightarrow k} \propto \prod_{j \in \partial i \setminus k} \bar{m}_{\sigma}^{j \rightarrow i}$$

$$\left(= \Pr_{\mu_{V \setminus i \rightarrow k}} [x_i = \sigma] \right)$$

we can then write marginals succinctly in terms of the messages:

$$\textcircled{\text{600}} \quad \Pr_{\mu} [x_i = \sigma] = m_{\sigma}^{i \rightarrow j} m_{\sigma}^{j \rightarrow i}$$

Combining $\textcircled{\text{I}}$ and $\textcircled{\text{II}}$ yields:

$$\textcircled{\text{rec}} \quad m_{\sigma}^{i \rightarrow k} \propto \prod_{j \in \partial i \setminus k} \sum_{s \in \mathcal{S} \setminus \{\sigma\}} m_s^{j \rightarrow i} \cdot \psi_{ij}(\sigma, s)$$

BP on trees:

- 1). Pick arbitrary root vertex
- 2). For every leaf j and parent i , initialize $m_{j \rightarrow i} = 1/2 \quad \forall \sigma \in \{\pm 1\}$
- 3). Use (rec) to compute \bar{m} 's via dynamic programming, starting from leaves
- 4). Use $\textcircled{\text{I}}$ to compute m 's
- 5). Use (***) to compute marginals

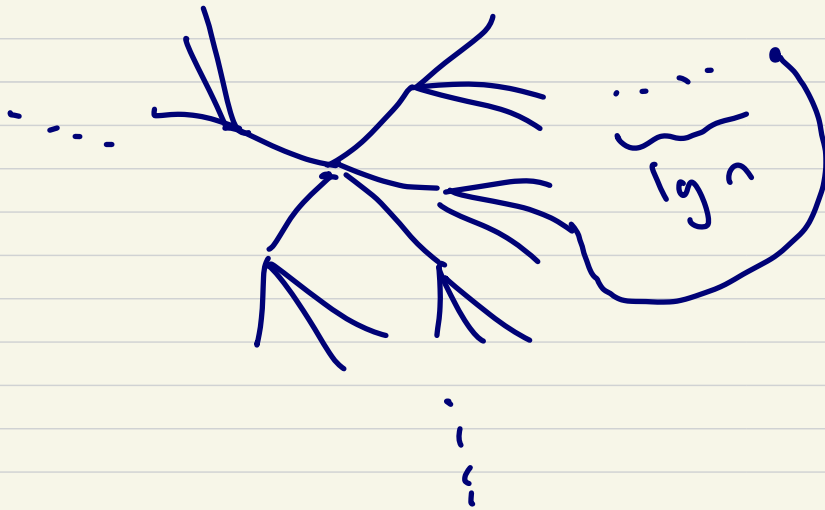
What if G is not a tree? Then subtree marginals $\{m_{j \rightarrow i}\}_{j \in \partial_i}$ are not independent...

Nevertheless, can still run the above algorithm* and hope it gives something interesting!

* As stated, the algorithm is stated w/ a tree structure in mind. Without this, we can still apply update rules for \bar{m} and m in parallel

over many rounds.

Intuition for why this is a good idea:
if the graph is a random sparse graph,
then locally it looks like a tree



If every edge appears w.p. $\frac{c}{n}$ for $c = O(1)$,
then probability that some "descendant" of depth d
"returns" to ancestor i is

$$1 - \left(1 - \frac{1}{n}\right)^{c^d}$$

So as long as $c^d \ll n$, this is $o(1)$.

Next lecture we will see a natural setting where
such a sparse random graph arises.

Even in such cases, BP is notoriously hard to analyze. We will instead see 2 rigorous alg's inspired by BP:

- 1). Spectral methods on nonbacktracking operators
- 2). approximate message passing.